

# An Introduction to the Fractional Brownian Motion Exposition & Insights

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- 1 Motivation of the Classical Process
- 2 Review of Gaussian Processes
- 3 The Fractional Brownian Motion & First Properties

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## Motivation

Consider a finite collection of i.i.d. random variables  $X_1; \dots; X_n$  with  $EX_i = 0$  and  $EX_i^2 = 1$ . Define a process  $Z_t /_{t \in [0;1]}$  by  $Z_{\frac{k}{n}} = X_k$  for  $k \in \{1; \dots; n\}$  and linearly interpolate on intervals of the form  $[\frac{k}{n}; \frac{k+1}{n}]$ .

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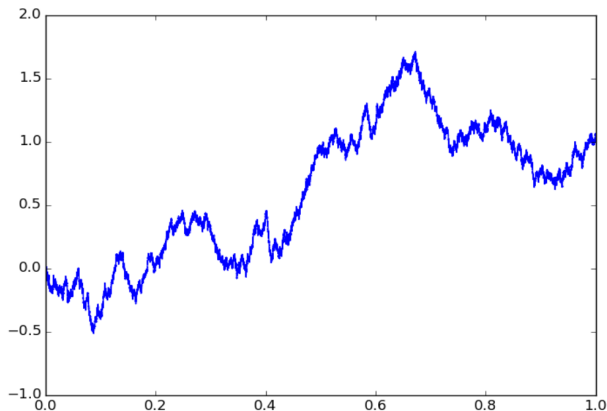
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# The Classical Process

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## Definition: Gaussian Process

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## Theorem: Characterization of Gaussian Processes

Gaussian processes are completely determined/characterized by their mean and covariance functions:

$$\begin{aligned} m : T \rightarrow \mathbb{R} & \quad ; \quad \quad : T \times T \rightarrow \mathbb{R} \\ t & \rightarrow E[X_t] & \quad ; \quad t, s & \rightarrow \text{Cov}(X_s, X_t) \end{aligned}$$

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i.e. Gaussian processes sharing the same mean and cov. functions are equal in law.

## Proof:

For a Gaussian process  $\{X_t\}_{t \in T}$  with mean and covariance functions  $m_X; \gamma_X$  and any choice  $t_1 < t_2 < \dots < t_p$  we have  $X = (X_{t_1}; \dots; X_{t_p})^T, t_i \in T$ , to be an  $\mathbb{R}^p$ -valued random variable with a multivariate normal distribution and hence to have characteristic function (Fourier transform of distribution):

$$\chi_X(u) = E[e^{i\langle u, X_{t_1}; \dots; X_{t_p} \rangle}] = \exp(i\langle u, m_X \rangle - \frac{1}{2} \langle u, \gamma_X u \rangle)$$

where  $m_X = (E[X_{t_1}]; \dots; E[X_{t_p}])^T = (m_X(t_j))_{1 \leq j \leq p}$  is the mean vector and  $\gamma_X = (\gamma_X(t_i, t_j))_{1 \leq i, j \leq p}$  the covariance matrix.

## Proof:

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$$\phi_X(u) = E[e^{i\langle u, X \rangle}] = \exp\left(i\langle u, \mu_X \rangle - \frac{1}{2} \langle u, \Sigma_X u \rangle\right);$$

where  $\mu_X = \{E[X_{t_i}]\}_{i=1}^p$  is the mean vector and  $\Sigma_X = \{\text{Cov}(X_{t_i}, X_{t_j})\}_{i,j=1}^p$  the covariance matrix. Now, considering another Gaussian process  $\{Y_t\}_{t \in T}$  with the same mean and covariance functions as  $\{X_t\}_{t \in T}$  i.e.  $\mu_Y = \mu_X$ , and  $\Sigma_Y = \Sigma_X$ , we have the characteristic function of its finite dimensional marginal,  $\{Y_{t_i}\}_{i=1}^p$ , to be given identically as to that of above.

## Proof:

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$$\phi_X(u) = E[e^{i\langle u, X \rangle}] = \exp\left(i\langle u, m_X \rangle - \frac{1}{2} \langle u, \Sigma_X u \rangle\right);$$

where  $m_X = \{E[X_{t_i}]\}_{i=1}^p$  is the mean vector and  $\Sigma_X = \{\Sigma_{X_{t_i}, X_{t_j}}\}_{i,j=1}^p$  the covariance matrix. Now, considering another Gaussian process  $\{Y_t\}_{t \in T}$  with the same mean and covariance functions as  $\{X_t\}_{t \in T}$  i.e.  $m_Y = m_X$ , and  $\Sigma_Y = \Sigma_X$ , we have the characteristic function of its finite dimensional marginal,  $\{Y_{t_i}\}_{i=1}^p$ , to be given identically as to that of above.

Hence, we have an agreement in distribution of the finite dimensional marginals of  $\{X_t\}_{t \in T}$  and  $\{Y_t\}_{t \in T}$  and hence an equivalence in law. □

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$$m : T \rightarrow \mathbb{R} \quad ; \quad \gamma : T \times T \rightarrow \mathbb{R}$$

$$m(t) = E[X_t] \quad ; \quad \gamma(s, t) = \text{Cov}(X_s, X_t)$$

i.e. Gaussian processes sharing the same mean and cov. functions are equal in law.

## Moral Conclusion:

Specifying Gaussian processes amounts to specifying a mean and covariance structure on the collection. Similarly as to how specifying a Gaussian r.v. requires only a mean and variance!



## Definition: Centered Gaussian Process

A **centered Gaussian process** is a stochastic process  $\{X_t\}_{t \in T}$  such that any finite linear combination of the variables  $X_t, t \in T$  is centered Gaussian. ( $t - E[X_t] = 0$ )

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## Definition: Brownian motion

A stochastic process  $\{B_t\}_{t \geq 0}$  on  $(\Omega, \mathcal{F}, P)$  is a standard **Brownian Motion** if it satisfies one of the following three equivalent assertions:

- (i)  $\{B_t\}_{t \geq 0}$  is a centered Gaussian process with covariance:  $\text{Cov}(B_s, B_t) = \min(s, t) \dots s \wedge t$ .
- (ii)  $B_0 = 0$  a.s. and for every  $0 \leq s < t$  the random variable  $B_t - B_s$  is independent of  $\{B_r : r \leq s\}$  and  $B_t - B_s \sim N(0, t - s)$ .
- (iii)  $B_0 = 0$  a.s. and for  $0 = t_0 < t_1 < \dots < t_n$  the r.v.'s  $B_{t_i} - B_{t_{i-1}}$  for  $1 \leq i \leq n$  are indep. and  $B_{t_i} - B_{t_{i-1}} \sim N(0, t_i - t_{i-1})$ .

and as well:

- (iv)  $\{B_t\}_{t \geq 0}$  has surely continuous sample paths:  $\forall \omega \in \Omega, t \mapsto B_t(\omega) \in C([0, \infty); \mathbb{R})$ .

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and as well:

- (iv)  $\{B_t\}_{t \geq 0}$  has surely continuous sample paths:  $\mathbb{P}\{\omega \in C : B_t(\omega) \text{ is continuous}\} = 1$ .

## Note:

Independence of increments are a key feature in the classical process!

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# The Fractional Brownian Motion

From here on out we consider only  $T = [0; \infty[$  i.e. positive time.

## Definition: Fractional Brownian Motion

A **fractional Brownian Motion** of *Hurst* parameter/index  $H \in ]0; 1[$  is a centered Gaussian process  $(B_t^H)_{t \geq 0}$  with covariance function

$$E[B_t^H B_s^H] = \frac{1}{2} (t^{2H} + s^{2H} - |t - s|^{2H});$$

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A fractional Brownian Motion of Hurst parameter/index  $H \in ]0; 1[$  is a centered Gaussian process  $B_t^H / t^{2H}$  with covariance function

$$E[B_t^H B_s^H] = \frac{1}{2} \cdot t^{2H} + s^{2H} - \frac{1}{2} |t - s|^{2H} /:$$

How is this a generalization of the classical Brownian motion?

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A fractional Brownian Motion of Hurst parameter/index  $H \in ]0; 1[$  is a centered Gaussian process  $B_t^H / \sqrt{t}$  with covariance function

$$E[B_t^H B_s^H] = \frac{1}{2} \cdot t^{2H} + s^{2H} - \frac{1}{2} |t - s|^{2H} /:$$

How is this a generalization of the classical Brownian motion?

If  $H = 1/2$  then fBm is nothing but a classical Brownian motion.

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## Proposition:

If  $H = 1/2$  then fBm is nothing but a classical Brownian motion.

## Proof:

Fix  $0 \leq s < t$  and observe

$$E[B_s^{1-2} B_t^{1-2}] = \frac{1}{2} \cdot t + s - \frac{1}{2} (t+s) = s = s \wedge t:$$





## Self-Similarity

For  $\alpha > 0$  given and  $H \in (0; 1/]$ , we have  $\alpha^{-H} B_{\alpha t}^H / \alpha^H$  is an fBm of Hurst index  $H$ .

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Illustration for  $H = 1/2$ :

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## Proof:

Exercise. □

## Self-Similarity

For  $\alpha > 0$  given and  $H \in (0; 1/]$ , the process  $\{\alpha^{-H} B_{t/\alpha}^H\}_{t \geq 0}$  is an fBm of Hurst index  $H$ .

## Stationary Increments

For  $s \geq 0$  fixed and  $H \in (0; 1/]$ , the process  $\{B_{s+t}^H - B_s^H\}_{t \geq 0}$  is an fBm of Hurst index  $H$ .

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For  $s, t > 0$  and  $H \in (0, 1/2]$ , the process  $B_{s+t}^H - B_s^H$  is an fBm of Hurst index  $H$ .

For  $\alpha > 0$  given and  $H \in (0; 1/2]$ , the process  $\alpha^{-H} B_t^H / t^{\alpha H}$  is an fBm of Hurst index  $H$ .

For  $s > 0$  fixed and  $H \in (0; 1/2]$ , the process  $B_{s+t}^H - B_s^H$  is an fBm of Hurst index  $H$ .



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## Stationary Increments

For  $s \geq 0$  fixed and  $H \in (0; 1/]$ , the process  $B_{s+t}^H - B_s^H /_{t_0}$  is an fBm of Hurst index  $H$ .

## Continuity of Sample Paths

An fBm  $B_t^H /_{t_0}$  admits a continuous modification. That is we have some process  $X_t /_{t_0}$  such that  $t \mapsto X_t \in C[0; \infty)$  (surely) and for all  $t \geq 0$ ,  $P(B_t^H = X_t) = 1$ .

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## Theorem: Kolmogorov-Čhenstov Continuity Theorem

Assume that for a stochastic process  $X_t$  there exists  $K > 0; p > 0; \alpha > 0$  such that for all  $s; t \geq 0$ :

$$E[|X_t - X_s|^p] \leq K |t - s|^{\alpha p}$$

Then the process has a continuous modification, i.e. a process  $X_t$  such that  $t \mapsto X_t \in C[0; \infty)$  and for all  $t \geq 0$   $P \cdot X_t = X_t = 1$ .

## Theorem: Kolmogorov-Čhenstov Continuity Theorem

Assume that for a stochastic process  $\{X_t\}_{t \geq 0}$  there exists  $K > 0; p > 0; \epsilon > 0$  such that for all  $s, t \geq 0$ :

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## Theorem: Kolmogorov-Čhenstov Continuity Theorem

Assume that for a stochastic process  $\{X_t\}_{t \geq 0}$  there exists  $K > 0; p > 0; \delta > 0$  such that for all  $s, t \geq 0$ :

$$E[|X_t - X_s|^p] \leq K \delta t * s^{\delta^{1+p}}$$

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## Proof:

Simply observe:

$$E[|B_t^H - B_s^H|^2] = \delta t * s^{2H};$$

and apply Kolmogorov-Čhenstov.



## Self-Similarity

For  $\alpha > 0$  given and  $H \in (0; 1/2]$ , the process  $\alpha^{-H} B_t^H / \alpha^{t/2}$  is an fBm of Hurst index  $H$ .

## Stationary Increments

For  $s \geq 0$  fixed and  $H \in (0; 1/2]$ , the process  $B_{s+t}^H - B_s^H$  is an fBm of Hurst index  $H$ .

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## The fBm is (in general) not Markov

Let  $B_t^H$  be a fractional Brownian motion of Hurst index  $H \neq 1/2$ . Then  $B_t^H$  is not a Markov process.

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Then how do we describe the dependence structure of fBm and how does such structure vary with the Hurst index?

## Dependence of Increments

Disjoint increments of an fBm of Hurst index  $H \in ]0; 1]$  are *negatively* correlated for  $H \in ]0; 1/2]$  and *positively* correlated for  $H \in ]1/2; 1/$ .



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Suppose that  $0 \leq s_1 < t_1 < s_2 < t_2$  so as to ensure  $[s_1; t_1] \cap [s_2; t_2] = \emptyset$ . Then one can check the covariance of the increments to be given as:

$$E[(B_{t_1}^H - B_{s_1}^H)(B_{t_2}^H - B_{s_2}^H)] = \frac{1}{2} (\alpha_2^H s_1^{2H} - \alpha_2^H t_1^{2H} - \alpha_2^H s_1^{2H} + \alpha_2^H t_2^{2H}) :$$

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$$E[B_{t_1}^H - B_{s_1}^H, B_{t_2}^H - B_{s_2}^H] = \frac{1}{2} (\alpha_2^{2H} s_1^{2H} - \alpha_2^{2H} t_1^{2H} - \alpha_2^{2H} s_1^{2H} + \alpha_2^{2H} t_2^{2H}) :$$

Then considering the map  $f(x) = x^{2H}$  and putting

$a_1 = t_2 - s_1; a_2 = t_2 - t_1; b_1 = s_2 - s_1; b_2 = s_2 - t_1$  gives that  $a_1 + a_2 = b_1 + b_2 = t_2 - s_1$  (note that  $b_2 < a_2 < b_1 < a_1$ ) and allows the above to be expressed as follows:

$$E[B_{t_1}^H - B_{s_1}^H, B_{t_2}^H - B_{s_2}^H] = \frac{1}{2} (f(a_1) - f(a_2) - f(b_1) + f(b_2)) :$$

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Suppose that  $0 \leq s_1 < t_1 < s_2 < t_2$  so as to ensure  $[s_1; t_1] \cap [s_2; t_2] = \emptyset$ . Then one can check the covariance of the increments to be given as:

$$E[B_{t_1}^H - B_{s_1}^H / B_{t_2}^H - B_{s_2}^H] = \frac{1}{2} \alpha_2^* s_1^{2H} - \alpha_2^* t_1^{2H} + \alpha_2^* s_1^{2H} - \alpha_2^* t_2^{2H} :$$

Then considering the map  $f(x) = x^{2H}$  and putting  $a_1 = t_2 - s_1; a_2 = t_2 - t_1; b_1 = s_2 - s_1; b_2 = s_2 - t_1$  gives that  $a_1 - a_2 = b_1 - b_2 = t_1 - s_1$  (note that  $b_2 < a_2 < b_1 < a_1$ ) and allows the above to be expressed as follows:

$$E[B_{t_1}^H - B_{s_1}^H / B_{t_2}^H - B_{s_2}^H] = \frac{1}{2} (f(a_1) - f(a_2) - f(b_1) + f(b_2)) :$$

Now as  $f'' < 0$  for  $H \in (0; 1/2]$  we have for such  $H$  that  $E[B_{t_1}^H - B_{s_1}^H / B_{t_2}^H - B_{s_2}^H] < 0$ .  
 And since  $f'' > 0$  for  $H \in (1/2; 1]$  we have for such  $H$  that  $E[B_{t_1}^H - B_{s_1}^H / B_{t_2}^H - B_{s_2}^H] > 0$ .

# Sample Path Regularity w.r.t $H$

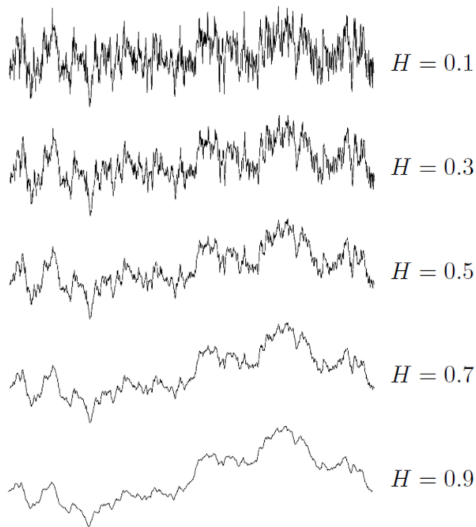


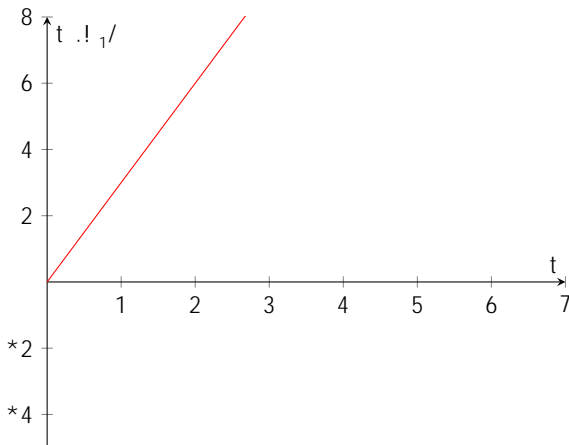
FIGURE 1. Paths of fBm for different values of  $H$ .

## Gaussian Beam; $H = 1$

One can check that for  $H = 1$  we have  $B_t^H / t^{d/2} = t / t_0$ , for  $t \in \mathbb{N} \cdot t_0$ .

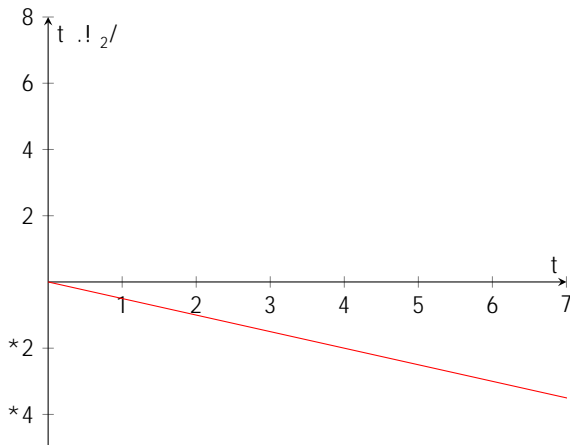
## Gaussian Beam; $H = 1$

One can check that for  $H = 1$  we have  $B_t^H / t_{g0} \stackrel{d}{=} t / t_{g0}$ , for  $t \in [0; 1]$ .



## Gaussian Beam; $H = 1$

One can check that for  $H = 1$  we have  $B_t^H / t^{H-d} = t / t_0$ , for  $t \in \mathbb{N} \setminus \{0\}$ .



## Gaussian Beam; $H = 1$

One can check that for  $H = 1$  we have  $B_t^H / t_{g0} \stackrel{d}{=} t / t_{g0}$ , for  $t \in \mathbb{N} \setminus \{0\}$ .

